# CONTAGION VERSUS FLIGHT TO QUALITY IN FINANCIAL MARKETS

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#### **Outline**

- Transmission of Risk between Economies
- Definitions of Interdependence and Contagion
- Statistical measures for dependence: Pitfalls of correlation
- Multivariate Extreme Value Theory: A new copula
- Measuring Interdependence and Contagion by tail dependence measures
- Causality in the Extremes
- Application: The flight to quality phenomenon

#### Transmission of Risk between Economies

Every economy is exposed to a series of factors that can culminate in what can be called crisis.

Types of crises: financial, liquidity, banking or currency crises.

**Definition 1.** A general definition of crisis in a market is given by a threshold level such that in case is exceeded, it results in the collapse of the system producing the triggering of negative effects in the rest of the markets.

**In summary:** A crisis in one market is characterized by the collapse not only of that market but by the negative effects produced on other markets.

Two ways of regarding dependence: (In particular in crises periods)

From the point of view of the direction (Causality in the Extremes).

From the point of view of the intensity: strength of the links in turmoil periods.

# Interdependence and Contagion

- Interdependence due to rational links between the variables (markets).
- Contagion effects: abnormal links between the markets triggered by some phenomena (crisis).

#### Regarding the direction:

- $\star$  Interdependence implies that both markets collapse because both are influenced by the same factors (Forbes and Rigobon (2001), Corsetti, Pericoli, Sbracia (2002)).
- \* Contagion implies that the collapse in one market produces the fall of the other market.

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#### Regarding the intensity:

- \* Interdependence implies no significant change in cross market relationships.
- \* Contagion implies that cross market linkages are stronger after a shock to one market.

# Transmission Channels connecting the markets

#### From an economic viewpoint:

• Economic fundamentals, market specific shocks, impact of bad news, phycological effects (herd behavior).

From an statistical viewpoint: Pearson correlation.

$$Corr(X_1, X_2) = \frac{E(X_1 - E(X_1))(X_2 - E(X_2))}{\sqrt{V(X_1)}\sqrt{V(X_2)}},$$

with  $X_1$  and  $X_2$  random variables.

Correlation is not sufficient to measure the dependence found in financial markets.

- It is only reliable when the random variables are jointly gaussian.
- Conditioning on extreme events can lead to misleading results.

#### **Pitfalls of Correlation**

These results are found in Embrechts, et al. (1999) and in Boyer et al. (1999).

- Correlation is an scalar measure (Not designed for the complete structure of dependence).
- A correlation of zero does not indicate independence between the variables.
- Correlation is not invariant under transformations of the risks.
- Correlation is only defined when the variances of the corresponding variables are finite.
- An increase in the correlation between two variables can be **JUST** due to an increase in the variance of one variable.

**Ex.**- Let  $\rho$  be the correlation between two r.v.'s X,Y and let us condition on  $X\in A$ . Then  $\rho_A=\rho\left(\rho^2+(1-\rho^2)\frac{V(X)}{V(X|A)}\right)^{-1/2}$ 

**SOLUTION:** A complete picture of the structure of dependence (Copula functions).

#### Copula functions for dependence

**Definition 2.** A function  $C:[0,1]^m \to [0,1]$  is a m-dimensional copula if it satisfies the following properties:

- (i) For all  $u_i \in [0,1]$ ,  $C(1,\ldots,1,u_i,1,\ldots,1) = u_i$ .
- (ii) For all  $u \in [0,1]^m$ ,  $C(u_1,\ldots,u_m)=0$  if at least one of the coordinates is zero.
- (iii) The volume of every box contained in  $[0,1]^m$  is non-negative, i.e.,  $V_C([u_1,\ldots,u_m]\times [v_1,\ldots,v_m])$  is non-negative. For m=2,  $V_C([u_1,u_2]\times [v_1,v_2])=C(u_2,v_2)-C(u_1,v_2)-C(u_2,v_1)+C(u_1,v_1)\geq 0$  for  $0\leq u_i,v_i\leq 1$ .

By Sklar's theorem (1959),

$$H(x_1, \ldots, x_m) = C(F_1(x_1), \ldots, F_m(x_m)),$$

with H the multivariate distribution, and  $F_i$  the margins.

#### Our Goal: Using dependence in the Extremes

Let  $(M_{n1}, \ldots, M_{nm})$  be the vector of maxima, and denote its distribution by

$$H^n(a_{n1}x_1+b_{n1},\ldots,a_{nm}x_m+b_{nm})=P\{a_{ni}^{-1}(M_{ni}-b_{ni})\leq x_i,i=1,\ldots,m\}.$$

The central result of EVT in the multivariate setting (mevt) is:

$$\lim_{n \to \infty} H^n(a_{n1}x_1 + b_{n1}, \dots, a_{nm}x_m + b_{nm}) = G(x_1, \dots, x_m),$$

with G a mevd.

**Theorem 1.** The class of mevd is precisely the class of max-stable distributions (Resnick (1987), proposition 5.9).

These distributions satisfy the following Invariance Property,

$$G^{t}(tx_{1},\ldots,tx_{m})=G(\alpha_{1}x_{1}+\beta_{1},\ldots,\alpha_{m}x_{m}+\beta_{m}),$$

for every t > 0, and some  $\alpha_j > 0$  and  $\beta_j$ .

By Sklar's theorem,

$$\lim_{n \to \infty} H^n(a_{n1}x_1 + b_{n1}, \dots, a_{nm}x_m + b_{nm}) = C(G_1(x_1), \dots, G_m(x_m)),$$

with  $G_i$  univariate evd.

Under an appropriate transformation of the margins  $(Z_i = 1/log \frac{1}{F_i(X)})$ ,

$$\lim_{n \to \infty} H^{*n}(nz_1, \dots, nz_m) = C(\Psi_1(z_1), \dots, \Psi_1(z_m)), \tag{1}$$

with  $\Psi_1(z) = \exp(-\frac{1}{z})$ , standard Fréchet, and the invariance property for copulas reads

$$C^{n}(\Psi_{1}(nz_{1}), \dots, \Psi_{1}(nz_{m})) = C(\Psi_{1}(z_{1}), \dots, \Psi_{1}(z_{m})).$$

Taking logs in both sides of (1) and applying the invariance property we have

$$\lim_{n \to \infty} \frac{H^*(nz_1, \dots, nz_m)}{1 + \log C(\Psi_1(nz_1), \dots, \Psi_1(nz_m))} = 1.$$

Then,  $H^*(z_1,\ldots,z_m)=C(\Psi_1(z_1),\ldots,\Psi_1(z_m))$ , from some threshold vector  $(z_1,\ldots,z_m)$  sufficiently high.

- ullet The copula function C is derived from the limiting distribution of the maximum.
- C must be of exponential type (extension of the EVT for the univariate case).

The Gumbel copula is within this class. Its general expression is

$$C_G(u_1, \dots, u_m; \theta) = \exp^{-[(-\log u_1)^{\theta} + \dots + (-\log u_m)^{\theta}]^{1/\theta}}, \quad \theta \ge 1,$$

with  $u_1, \ldots, u_m \in [0, 1]$  and  $\theta \geq 1$ .

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**Inconvenient:** This multivariate extreme value distribution describes the dependence between the variables for the  $multivariate\ upper\ tail\ ((z_1,\ldots,z_m)\ sufficiently\ high).$ 

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**Our aim:** Modelling the complete structure of dependence between the variables. Not just the relation in the extremes!

### Our Contribution: A NEW Copula

WE PROPOSE instead (for m=2):

$$\widetilde{C}_G(u_1, u_2; \Theta) = \exp^{-D(u_1, u_2; \gamma, \eta)[(-\log u_1)^{\theta} + (-\log u_2)^{\theta}]^{1/\theta}}, \tag{2}$$

with

$$D(u_1, u_2; \gamma, \eta) = \exp^{\gamma(1-u_1)(1-u_2)^{\eta}}, \quad \gamma \ge 0, \quad \eta > 0.$$
(3)

The function  $D(u_1,u_2;\gamma,\eta)$  accommodates departures from the invariance property with  $\gamma>0$  and  $\eta\neq 1$ .

**Theorem 2.** The function  $\widetilde{C}_G: [0,1] \times [0,1] \to [0,1]$  defined in (2) and (3) is a copula function if the parameters in  $\Theta$  satisfy that  $\widetilde{c}_G(u_1,u_2;\Theta) > 0$ ,  $\forall (u_1,u_2) \in [0,1] \times [0,1]$ , with  $\widetilde{c}_G(u_1,u_2;\Theta) = \frac{\delta^2 \widetilde{C}_G(u_1,u_2;\Theta)}{du_1 du_2}$  the density function of the copula  $\widetilde{C}_G$ .

## Advantages of this NEW Copula

- This copula function is derived from the multivariate extreme value theory, in contrast to ad-hoc models for the dependence structure.
- The function  $D(u_1, u_2; \gamma, \eta)$  and in particular the parameter  $\gamma$  extend the multivariate extreme value theory results to the entire range of the random variables.
- $\widetilde{C}_G$  is able to explain asymmetric effects of the variables for  $\eta \neq 1$ , and  $\gamma > 0$ .
- This copula is sufficiently flexible to describe different forms of dependence,
  - $\star$  Dependence:  $\theta \neq 1$  or  $\theta = 1$  and  $\gamma > 0$ .
  - $\star$  Independence:  $\gamma = 0$ ,  $\theta = 1$ .
  - $\star$  Asymptotic dependence:  $\theta > 1$ .
  - $\star$  Asymptotic independence:  $\theta = 1$ .

### Our Contribution: Tail Dependence Measures

Alternatives to the standard ℵ,

$$\aleph = \lim_{t \to \infty} P\{Z_2 > t | Z_1 > t\},$$

introduced by Ledford and Tawn (1997) and Coles, Heffernan and Tawn (1999).

- Definitions of Interdependence and Contagion by means of tail dependence measures.
- The translation of these definitions to mathematical expressions by using copula functions.
- The distinction between types of contagion: In Intensity and In the direction.

#### Interdependence

Lehman (1966) defined two random variables  $Z_1, Z_2$  as positively quadrant dependent (PQD) if for all  $(z_1, z_2) \in \mathbb{R}^2$ ,

$$P\{Z_1 > z_1, Z_2 > z_2\} \ge P\{Z_1 > z_1\}P\{Z_2 > z_2\},$$

or equivalently if

$$P\{Z_1 \le z_1, Z_2 \le z_2\} \ge P\{Z_1 \le z_1\}P\{Z_2 \le z_2\}.$$

**Definition 3.** Two random variables are Interdependent if they are PQD. Interdependence is characterized by joint movements in the same direction (co-movements).

In terms of the copula Interdependence amounts to see that  $g(u_1, u_2) > 0$ , with

$$g(u_1, u_2) = \widetilde{C}_G(u_1, u_2) - u_1 u_2.$$

# **Contagion in Intensity**

A stronger condition is required: **Tail Monotonicity**.

**Definition 4.** Suppose  $Z_1$ ,  $Z_2$  with common  $\Psi_1$  and consider z a threshold that determines the extremes in the upper tail of both random variables. There exists a contagion effect between  $Z_1$  and  $Z_2$  if  $g(u_1, u_2)$  is an increasing function for both random variables, and for  $u_1, u_2 \geq u$  with  $u = \Psi_1(z)$ .

For the lower tails contagion in intensity is characterized by decreasing tail monotonicity in

$$P\{Z_1 \le z_1, Z_2 \le z_2\} - P\{Z_1 \le z_1\}P\{Z_2 \le z_2\}.$$

In terms of copulas contagion in the upper tails amounts to

$$h_1(u_1, u_2) = \frac{\delta \widetilde{C}_G(u_1, u_2)}{du_1} - u_2 > 0, \qquad h_2(u_1, u_2) = \frac{\delta \widetilde{C}_G(u_1, u_2)}{du_2} - u_1 > 0.$$

## Directional Contagion: Causality in the Extremes

The conditional probability is interpreted as a causality relationship.

Let z be a threshold determining the extremes for both random variables.

**Motivation:**  $P\{Z_2 > z' | Z_1 > z\} > P\{Z_2 > z'\} \stackrel{?}{\equiv} Z_1 \Rightarrow Z_2$ , with z' > z.

( $Z_1$  taking on extreme values is causing that  $Z_2$  takes on extreme values).

However, This is not true!

#### **False Intuition:**

$$P\{Z_2 > z' | Z_1 > z\} > P\{Z_2 > z'\} \equiv P\{Z_2 > z', Z_1 > z\} > P\{Z_2 > z'\}P\{Z_1 > z\}$$

This condition determines Contagion in Intensity NOT in the direction (No causality).

Assuming a common marginal d.f.  $\Psi_1$ , and a threshold z determining the extremes for both random variables, we find contagion spill-over from  $Z_1$  to  $Z_2$  if

$$P\{Z_2 > z' | Z_1 > z\} > P\{Z_1 > z' | Z_2 > z\},$$

or equivalently if

$$P\{Z_2 \leq z' | Z_1 \leq z\} > P\{Z_1 \leq z' | Z_2 \leq z\}.$$

These conditions boil down to see

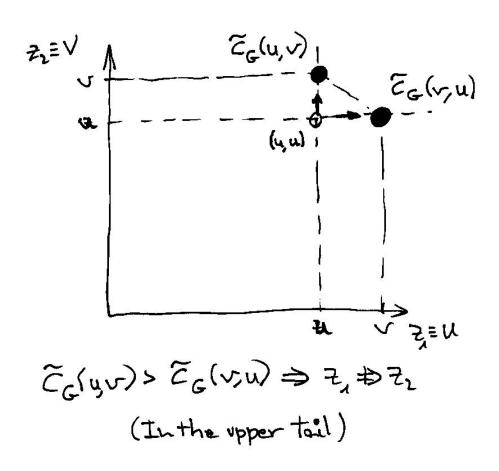
$$\widetilde{C}_G(u,v) > \widetilde{C}_G(v,u)$$
 for  $Z_1 \Rightarrow Z_2$  (Causality in the extremes),

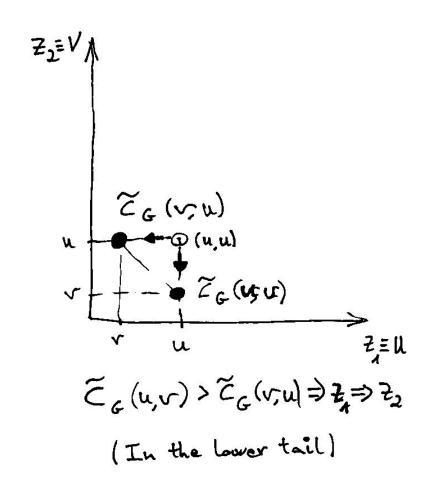
with  $u = \Psi_1(z)$ , and  $v = \Psi_1(z')$ .

Define  $gd_v(u) = \widetilde{C}_G(u,v) - \widetilde{C}_G(v,u)$ . Then

**Definition 5.**  $Z_1$  is influencing  $Z_2$  in the extreme values (contagion effect) if  $gd_v(u)$  is strictly positive for all v > u for the upper tail, and for all v < u for the lower tail, with  $u = \Psi_1(z)$ .

#### Intuition





# Application: Flight to quality versus Contagion

**Definition 6.** Outflows of capital from the stock markets  $(Z_2)$  to the bonds markets  $(Z_1)$  in crises periods.

This is represented by

$$P\{Z_1 > z | Z_2 < z'\} - P\{Z_1 > z\} > 0,$$

with z defining the extreme values in the upper tail, and z' in the lower tail.

**Experiment:** Dow Jones Corporate 02 Years Bond Index (DJBI02) vs Dow Jones Industrial Average: Dow 30 Industrial Stock Price Index (DJSI).

#### **General Model:**

$$X_{1,t} = g_1(X_{1,t-1}, X_{2,t-1}) + \varepsilon_{1,t}$$

$$X_{2,t} = g_2(X_{1,t-1}, X_{2,t-1}) + \varepsilon_{2,t}$$

with  $(\varepsilon_{1,t},\varepsilon_{2,t})\sim C_G$ .

**Financial Sequence:**  $X_{i,t} = 100 \ (log P_{i,t} - log P_{i,t-1})$ , i = 1, 2, with  $P_{i,t}$  the corresponding prices.

#### **Modelling Rational Dependence**

DJBI02 index is well modelled by an AR(1)-GARCH(1,1) model as follows,

$$X_{1,t} = 0.00025 + 0.089X_{1,t-1} + \sigma_{1,t}\varepsilon_{1,t}$$
, with  $\varepsilon_{1,t}$  i.i.d.  $(0,1)$ ,

and 
$$\sigma_{1,t}^2 = 6.194 \cdot 10^{-8} + 0.071 \varepsilon_{1,t-1}^2 + 0.903 \sigma_{1,t-1}^2$$
.

DJSI Index is modelled by the following pure GARCH(1,1) model,

$$X_{2,t}=\sigma_{2,t}arepsilon_{2,t}$$
 , with  $arepsilon_{2,t}$  i.i.d.  $(0,1)$ ,

and 
$$\sigma_{2,t}^2 = 3.0012 \cdot 10^{-6} + 0.096 \varepsilon_{2,t-1}^2 + 0.887 \sigma_{2,t-1}^2$$
.

The evolution of prices in one market is independent of the other.

The irrational dependence (dependence in the innovations) is measured by the links between the vectors  $(\varepsilon_{1,t}, \varepsilon_{2,t})$  and  $\widetilde{C}_G$ .

Estimate of 
$$\widetilde{C}_G$$
:  $\hat{\theta}_n = 1.031$ ,  $\hat{\eta}_n = 1$  and  $\hat{\gamma}_n = 0.175$ . ( $\Downarrow$ )

#### **IGARCH Effect**

Consider

$$X_t = \sigma_t \varepsilon_t$$
, with  $\varepsilon_t$  i.i.d.  $(0,1)$ ,

and

$$\sigma_t^2 = \omega + \alpha \varepsilon_{t-1}^2 + \beta \sigma_{t-1}^2,$$

with  $\alpha + \beta = 1$ .

#### Features of the model:

- $V(X_t) = \infty$ .
- In the same way that I(1) represents **persistence** in linear models, IGARCH(1,1) describes **persistence** in the square and absolute observations.
- Persistence, NOT Long Range Dependence, because the latter implies finite marginal variances.

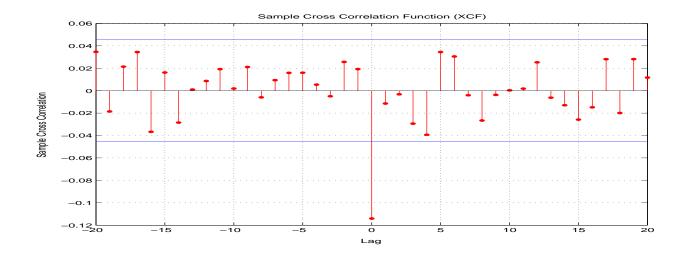
However, the IGARCH effect may show up by (Mikosch and  $St\breve{a}ric\breve{a}$ ):

- Persistence in the squares (true IGARCH).
- Non-stationarity due to different regimes (different means, different stationary GARCH, etc.)

#### Regarding Contagion:

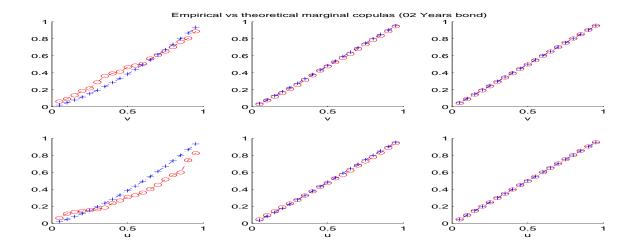
- For true IGARCH: Study the contagion effect for the vector of innovations  $(\varepsilon_1, \varepsilon_2)$  obtained from the IGARCH model.
- Non-stationarity: Consider the univariate sequence  $\{X_t\}$  and filter it by the corresponding regimes to obtain a sequence of innovations  $\varepsilon_t$  that is I(0) and serially independent.

# **Modelling Irrational Dependence**



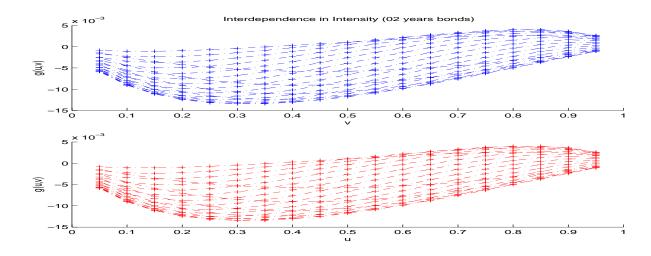
Cross correlation for different lags of the bivariate innovation sequence, spanning the period 02/01/1997 - 24/09/2004, n=1942 observations.

#### **Goodness of Fit Test**



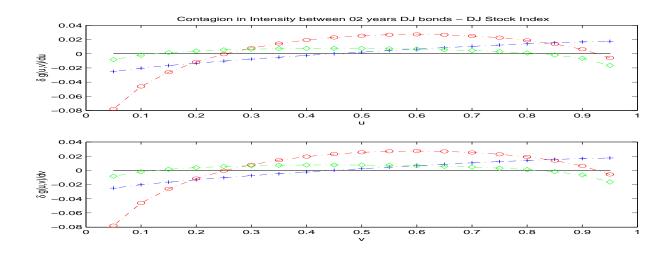
Empirical (o-) and theoretical (+-) margins. The upper panel for the vertical sections and the lower panel for the horizontal section. 0.05 quantile, 0.50 quantile and 0.95 quantile respectively.

# Interdependence in Intensity



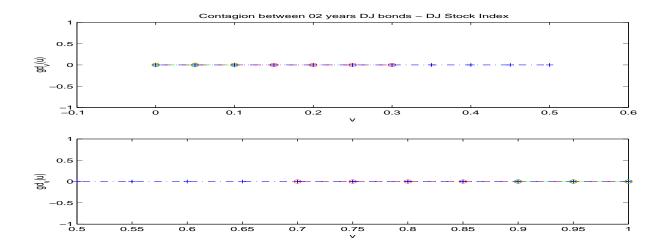
The upper panel depicts the function g(u, v) plotted against the innovations of DJSI. The lower panel g(u, v) plotted against the innovations of DJBI02.

# **Contagion in Intensity**



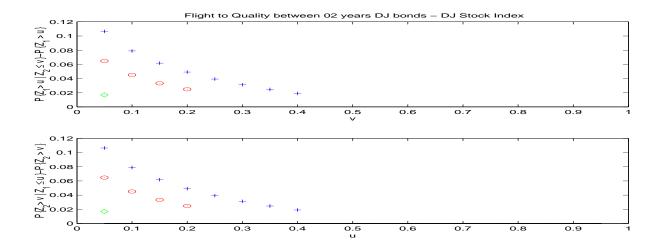
The upper panel depicts  $h_1(u,v)$  against DJBI02 and the lower panel depicts  $h_2(u,v)$  against DJSI. (+-) for 0.05 quantile, (o-) the 0.50 quantile and  $(\diamond-)$  the 0.95 quantile.

#### **Directional Contagion**



The upper panel depicts  $gd_v(u)$  for  $v \leq u$ . (+-) for u=0.50, (o-) for u=0.30 and  $(\diamond-)$  for u=0.10. The lower panel depicts  $gd_v(u)$  for v>u. (+-) for the u=0.50, (o-) for u=0.70 and  $(\diamond-)$  for u=0.90.

# Flight to Quality: $P\{Z_1 > u | Z_2 < v\} - P\{Z_1 > u\} > 0$



In the upper panel (+-) for u=0.60, (o-) for u=0.80 and  $(\diamond-)$  for u=0.95. In the lower panel (+-) for v=0.60, (o-) for v=0.80 and  $(\diamond-)$  for v=0.95.

# **Some Interesting Facts**

- Negative interdependence in the left tail, that turns positive in the right tail.
- Absence of directional contagion (Symmetric effects between the variables).
- Strong opposite movements in the middle of the domain (negative interdependence) that decrease when the variables take on extreme values. (Intensity Contagion without Interdependence).
- Evidence of Flight to Quality in both tails. This phenomenon can be interpreted as a substitution effect between bonds (DJBI02) and stocks (DJSI) when either of the sequences are in crises periods.
- DJBI02 depends on its past and in the volatility dynamics.
- DJSI depends only on its volatility dynamics.